### DOCUMENT RESUME

ED 410 262 TM 027 038

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TITLE Do Metacognitive Skills and Learning Strategies Transfer

across Domains?

SPONS AGENCY

College Entrance Examination Board, New York, N.Y.

PUB DATE

Mar 97

NOTE

23p.; Paper presented at the Annual Meeting of the American Educational Research Association (Chicago, IL, March 24-28,

1997).

PUB TYPE

Reports - Research (143) -- Speeches/Meeting Papers (150)

EDRS PRICE

MF01/PC01 Plus Postage.

DESCRIPTORS

Higher Education; Knowledge Level; \*Learning Strategies;
\*Mathematics Instruction; \*Metacognition; Models; Theory
Practice Relationship; \*Transfer of Training; Undergraduate

Students; \*Verbal Ability

**IDENTIFIERS** 

Domain Knowledge; \*Monitoring; Self Regulated Learning

### ABSTRACT

Current theories of metacognition suggest that effective control of learning by either metacognitive or self-regulatory processes cannot occur without accurate monitoring of learning. Given this theoretical framework, there are questions of whether knowledge monitoring and self-regulated learning abilities are domain-specific or whether metacognitive processes, in particular knowledge monitoring ability, generalize across academic domains. This study examines that issue by exploring the correlations among measures of metacognitive knowledge, learning, and study strategies, and academic achievement across the domains of verbal ability and mathematics. Using parallel measures of knowledge monitoring in both the verbal and mathematical domains, 120 undergraduates estimated their metacognitive knowledge, reported their Vinsidence in the accuracy of those estimates, and completed a self-report measure of learning and study strategies. Results suggest that metacognitive knowledge is generalizable across both the verbal and mathematical domains. The correlations between the two knowledge monitoring measures and students' confidence estimates were also in the expected directions. Moreover, both knowledge monitoring measures correlated with students' grade point averages. Correlations with subscales of the Learning and Study Strategies Inventory were not significant. Findings are discussed in the framework of current theory in metacognition and conceptions of strategic learning. An appendix shows multiple regression results. (Contains 1 table and 47 references.) (Author/SLD)

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# Do Metacognitive Skills and Learning Strategies Transfer Across Domains?

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This paper was presented at a symposium on Assessing Metacognitive Knowledge

Monitoring held at the annual convention of the American Educational Research

Association, Chicago, March 1997. The research was supported by the College Board.



### **ABSTRACT**

Current theories of metacognition suggest that effective control of learning by either metacognitive or self-regulatory processes cannot occur without accurate monitoring of learning. Given this theoretical framework, there are questions of whether knowledge monitoring and self-regulated learning abilities are domain specific or whether metacognitive processes, in particular knowledge monitoring ability, generalizes across academic domains. This study examines that issue by exploring the correlations among measures of metacognitive knowledge monitoring, learning and study strategies, and academic achievement across the domains of verbal ability and mathematics. Using parallel measures of knowledge monitoring in both the verbal and mathematical domains, 120 students estimated their metacognitive knowledge, reported their confidence in the accuracy of those estimates, and completed a self-report measure of learning and study strategies. Results suggest that metacognitive knowledge monitoring is generalizable across both the verbal and mathematical domains (r = .49, p < .01). The correlations between the two knowledge monitoring measures and students' confidence estimates (r=.30, and .53, p<.01) were also in the expected directions. Moreover, both knowledge monitoring measures correlated with students' gpa (r = .35 and .24, p < .01, respectively). The correlations with the LASSI subscales, however, were not significant, although a number of LASSI subscale scores did correlate significantly with gpa. The findings are discussed within the framework of current theory in metacognition and contemporary conceptions of strategic learning.



### INTRODUCTION

Metacognition has been defined as the ability to monitor, evaluate, and make plans for one's learning (Flavell, 1979; Brown 1980). The literature in this area identifies two distinct aspects of metacognition, knowledge about cognition and the regulation of cognition, with both viewed as important for effective learning (Brown, 1987; Garner and Alexander, 1989; Jacobs and Paris, 1987). Indeed, researchers have reported differences in metacognitive abilities between capable and less capable learners (see, for example, Baker, 1989; Brown and Campione, 1986; Garner and Alexander, 1989; Pressley and Ghatala, 1990). In general, students with effective metacognitive skills accurately estimate their knowledge in a variety of domains, monitor their on-going learning, update their knowledge, and develop effective plans for new learning.

Though widely recognized as important, assessing individual differences in metacognition has proven to be both difficult and time consuming (O'Neil, 1991; Schwartz and Metcalfe, 1994), and remains an obstacle to the advance of research. Typically, assessments of metacognition rely either on inferences from classroom performance, or ratings based on interviews of students who are questioned about their knowledge and cognitive processing strategies, or on analyses of "think-aloud" protocols (Meichenbaum, Burland, Gruson, & Cameron, 1985). Recently, a number of self-report measures of metacognition (Everson, Hartman, Tobias, and Gourgey, 1991; O'Neil, 1991; Pintrich, Smith, Garcia, and McKeachie, 1991; Schraw and Dennison, 1994) have been developed. For the most part, these measures are more efficiently administered and scored than "think aloud" protocols. Unfortunately, the use of self-report measures raises questions of validity (see Schwartz and Metcalfe (1994) for a review of these methodological issues). In light of these concerns, it is not surprising that little research has been conducted on the metacognitive processes related to learning in adults, looking, for example, at those in college or in advanced instructional or training programs, where instructional times less easily accommodates research. Thus, more efficient measures of metacognition are needed not merely to satisfy psychometric standards (although important), but because they would permit research in settings where instructional time is less flexible, such as college classrooms and training courses.



In this paper we introduce a method for assessing students' knowledge monitoring abilities (referred to generally as the KMA) in two salient academic domains, verbal ability and mathematics, and we relate these measures to self-report indices of learning and study strategies and to academic performance in college. Before presenting our results, it may be useful to establish the context for investigating the relationship between metacognition and complex learning in environments such as college and industry-based training courses.

### Metacognition and Learning

In college students learn a great deal of new knowledge, and are faced, at times, with classroom and laboratory situations that require them to learn material and apply problem solving skills in new and innovative ways. The literature on human metacognition makes a compelling case for its importance in these learning and training environments (Bjork, 1994; Davidson, Deuser, and Sternberg, 1994). Accurate monitoring of new learning enables students with effective metacognitive strategies to concentrate on new content and adjust their learning goals. In college classrooms or advanced training programs, for example, the learner usually has to master a great deal of new knowledge in a limited amount of time. Moreover, learning in classrooms or other structured training environments is often dynamic, with knowledge and information being acquired and updated frequently. Clearly, those who accurately distinguish between what they have already mastered and what is yet to be learned have an advantage in these situations, since they can be more strategic and effective learners. Yet many students have ineffective metacognitive strategies. It is important, therefore, to evaluate students' metacognitive abilities and target instruction to the development of these key learning strategies.

### Monitoring Knowledge

Given the premise outlined above, we assumed that knowledge monitoring accuracy, an ability presumably involved in the regulation of cognition, would be related to learning in complex environments and reflected in indices such as students' grades in college. Thus, we developed a technique for assessing this metacognitive dimension that conjointly evaluates students' self-reports of their knowledge in a domain (e.g., verbal ability or mathematics) and their performance on an objective measure of knowledge in that domain (see, for example, Everson, Smodlaka, and Tobias, 1994; Tobias and Everson, 1996;



Tobias and Everson, in press). The basic strategy is to assess knowledge monitoring by evaluating the differences between students' estimates of their knowledge in a particular domain (both procedural and declarative) and their actual knowledge as determined by performance on a test. In the prototypical KMA, students are asked to estimate their knowledge (e.g., in the verbal domain they identify words they know or do not know from a word list, or in mathematics its in problems they expect they can solve) and these estimates are contrasted with their performance on a standardized test containing many of the same words or math problems. Differences between students' estimates and their test performance provide an index of knowledge monitoring ability. This approach is similar to methods used in research on metamemory (Nelson and Narens, 1990), reading comprehension (Glenberg, Sanocki, Epstein and Morris, 1987), and psychophysics (Green and Swets, 1966). A brief description of our use of the KMA in an earlier study (Everson et al., 1994) serves as an illustration.

In an effort to understand better the relationship between metacognition and reading comprehension, the KMA was administered to 169 college students. Each was given a list of 33 words and asked to indicate the words they knew and did not know. This was followed by a vocabulary test based on the same words. The KMA generated four scores, including estimates that the word was: a) known and correctly identified on a subsequent vocabulary test [++]; b) known, yet incorrectly identified on the test [+-]; c) unknown, yet correctly identified on the test [-+]; and d) unknown and incorrectly identified on the test [--]. Within this framework the [++] and the [--] scores represented accurate metacognitive estimates of vocabulary word knowledge, while the two other measures [+-, and -+] represented inaccurate knowledge monitoring estimates. The results indicated that college students' accurate metacognitive judgments, both the + + and - - scores, were positively correlated with their scores on a standardized measure of reading comprehension (i.e., the Descriptive Test of Language Skills, 1979), r = .46 and -.43, respectively. Encouraged by these findings, we adapted the KMA for use in an extensive program of research (Tobias and Everson, 1996; Tobias and Everson, in press).

### Learning and Study Strategies

As we noted earlier, there has been a growing research interest in the role that metacognitive abilities play in facilitating learning in complex environments. In addition to



the work on metacognitive knowledge monitoring, interest in the more general learning-to-learn phenomena continues to increase (see, for example, Nickerson, Perkins, & Smith, 1985; Weinstein & Mayer, 1986; Weinstein, Goetz, & Alexander, 1988; Zimmerman & Schunk, 1989). In general, this research suggests that effective learning depends not only on what the learner knows about a given subject or academic domain, but also what the learner thinks about prior to, during, and after the learning activity—i.e., the strategic aspects of learning. Weinstein (Weinstein & Mayer, 1986; Weinstein et al., 1988) refers to these as learning and study strategies, which include... "a variety of cognitive processes and behavioral skills designed to enhance learning effectiveness and efficiency" (Weinstein & Meyer, 1991, p. 41). In an effort to assess these aspects of effective learning, Weinstein, Schulte, and Palmer (1987) developed the Learning and Study Strategies Inventory (LASSI), a self-report instrument that includes measures of attitude, motivation, time management, anxiety, concentration, information processing, selecting main ideas, study aids, self testing, and test strategies.

In an effort to refine our understanding of the cognitive processes underlying effective learning, we sought to examine the patterns of relationships among the LASSI subscales to the knowledge monitoring estimates derived from the KMA measures in both the verbal and mathematical domains. By viewing the metacognitive components of learning in this manner, we hope to gain additional understanding of the contextual boundaries of cognition (Perkins & Salomon, 1989). Thus, the objectives of the research reported below are (1) to examine the empirical relationships between and among the KMA scores, the LASSI subscale scores, and indices of learning in college; and (2) to address the question of whether the metacognitive knowledge monitoring skills are generalizable across academic domains.

### *METHOD*

### **Participants**

A total of 120 undergraduates from an urban college participated in the study, 73 females and 47 males. Participants were recruited from an introductory psychology course, and were awarded extra course credit for taking part in the study.



### Materials and Procedures

Experimental materials in this study were administered in three distinct large-scale (N=40) testing sessions. At the beginning of each session participants were administered a paper-and-pencil version of the Learning and Study Strategies Inventory (LASSI). The LASSI is a 77-item self-report instrument that includes ten subscales: Attitude, Motivation, Time Management, Anxiety, Concentration, Information Processing, Selecting Main Ideas, Study Aids, Self Testing, and Test Strategies. The individual items on each scale are presented using a 5 point Likert-type scale. The LASSI scales have reported reliabilities that range from a = .68 to .86 (Weinstein et al., 1987). After completing the LASSI, the KMA measures were administered to participants in a group setting.

Two KMA measures were administered: the first, the KMA verbal (KMAV), included the presentation of 39 words adopted from the Nelson-Denny Vocabulary subscale (Riverside Press, 1979) appropriate for post-secondary students, with difficulty estimates ranging from easy (p = .84) to hard (p = .43); the second KMA, mathematics (KMAM), included the presentation of 21 math problems adapted from the Preliminary Scholastic Assessment Test (College Board, 1994). Each of the KMA items was presented singly to the participant groups using a Performa computer projection device. The KMA math items were programmed to remain visible for a delay of 20 seconds, the vocabulary words were presented for 8 seconds. The proper delay was determined for each of the two sets of items in an earlier pilot study. The delay was designed to afford the participants only enough time to read the problem, quickly determine if they could solve it, and indicate this on a response form.

Participants were instructed to view the math problems and vocabulary items and assess whether they could or could not answer the item correctly, if given enough time. They were then asked to indicate, separately, the degree of confidence they had in that decision. The degree of confidence was presented in a 5 point Likert type scale ranging from 0 to 100% confident. Following the estimates, participants were asked to complete a form indicating their current and projected grades in Mathematics and English subject areas. Participants were then administered the paper-and-pencil multiple-choice math and vocabulary tests which contained the identical sets of items presented earlier. The 21 item math test and the 39 item vocabulary test were scored as correct or incorrect, and the



Hamman coefficient (Schraw, 1995), a measure of the accuracy of the estimates, was computed for each participant.

### RESULTS

Following suggestions made by Schraw (1995), the KMA math and vocabulary scoring methods discussed earlier were transformed from simple measures of association or recognition to a measure of the accuracy of the knowledge monitoring estimates. The Hamman coefficient (HC) described by Romesberg (1984) and subsequently discussed by Schraw (1995), was used to transform the KMAV and KMAM scores into the appropriate metric for subsequent analyses. In general, the HC, and by extension the KMAV and KMAM, varies from -1 to 1, with higher values indicating more accurate estimates of knowledge. Both the KMAV and KMAM scores were related to students' academic achievement using their undergraduate gpa's, as well as to scores derived from the LASSI subscales. The zero-order correlations of the two KMA scores, the confidence estimates, the LASSI subscales, and gpa are presented in Table 1.

### **INSERT TABLE 1 HERE**

The correlation between the two KMA measures, math (KMAM) and word knowledge (KMAV) was significant, r = .49 (p < .01), suggesting that knowledge monitoring abilities tend to generalize, albeit not very powerfully, across academic domains. The estimates of students' confidence in their knowledge monitoring abilities with reference to test response were also moderately correlated with metacognitive performance in the math (r = .53, p< .01) and verbal (r = .30, p < .01) domains. Students' confidence in their estimates were more weakly correlated across the math and verbal domains (r = .2653, p< .01), suggesting that domain specificity may be more salient for this type of confidence measure.

In Table 1 there are twenty zero-order correlations reported for the LASSI subscales (N=10) and the two KMA measures. Of the twenty, only four were significant and, somewhat unexpectedly, they were all negative. The correlations, for example, between KMAM and LASSI measures of motivation, mental self-testing, and time management



skills (r = -.24, -.21, and -.23 (p < .05), respectively), as well as the correlation between KMAV and mental self-testing (r = -.19 (p < .05), indicate that students' self-reports of poor strategy use and undeveloped skills in these areas was not aligned with their metacognitive knowledge monitoring as measured more objectively with the KMA procedure. Interestingly, both the two KMA scores and the LASSI subscale scores were, for the most part, significantly correlated with academic achievement as measured by gpa.

Multiple regression methods were used to explore further students' KMA scores, as well as the complex relationship of their KMA scores, LASSI self-reports, and academic achievement. Looking first at the math KMA scores, the results of the regression analyses suggest that both the students' confidence estimates and their self-reports of academic motivation explain a significant portion of the variance ( $R^2 = .33$ , F(2, 113) = 27.94, p < .001). With respect to the vocabulary word knowledge KMA scores, the regression analysis reveals that students' confidence estimates and their self-reports of the self-testing strategies explain a small, but significant, proportion of the variance ( $R^2 = .21$ , F(2, 109) = 14.78, P < .001). As noted earlier, in both of these exploratory measurement models the LASSI self-report measures were negatively related to performance on the KMA's.

In order to understand the joint influences of the KMA scores and LASSI self-reports on academic success, multiple regression analysis was used in much the same way. The dependent variable in these analyses, however, was the total gpa as reported by the college. The regression analyses suggested that the KMA verbal score, as well as the LASSI subscale measuring self testing and reviewing strategies, explained a portion of the variance in students' gpa variance ( $R^2 = .23$ , F(2, 104) = 15.91, p < .001). It is important to note that the math KMA score did not enter the regression equation, and that the LASSI test-taking subscore, unlike the other LASSI scores discussed earlier, was positively related to academic achievement. (See Appendix A for a more complete presentation of the regression analyses.)

### DISCUSSION

Knowledge monitoring is an important component of academic success. The findings of this study, along with findings of a series of studies in this area (Everson and Tobias, in press; Tobias and Everson, in press) support this general conclusion. Moreover, this study



continues a line of research that supports the validity of the KMA, both in terms of the construct of metacognitive knowledge monitoring and the predictive validity of the assessment procedure. Interestingly, the results of this study suggest that knowledge monitoring may have a relatively large domain specific component, and that a general metacognitive knowledge monitoring ability may be context-bound. This argues, we suspect, for developing KMA procedures that are more domain specific, e.g., in science and social science. On the other hand, the vocabulary KMA scores more successfully differentiated the academically capable students than did the math KMA scores in this study, partially replicating findings reported elsewhere (Tobias and Everson, 1996). Also, the correlation between the verbal KMA score and gpa was higher than for the math KMA. This suggest that verbal metacognitive knowledge monitoring scores may be more salient predictors of college, since effective learning may be largely related to students' ability to process text-based materials across a variety of academic domains. Given the large body of research indicating that vocabulary test scores are one of the most powerful predictors of school learning (Breland, Jones, & Jenkins, 1994; Just & Carpenter, 1987), this finding is not altogether surprising.

Learning in complex domains such as science and engineering, or making diagnoses in medicine or other fields, often requires that students bring substantial amounts of prior learning to bear in order to understand and acquire new knowledge or solve problems. Some prior learning may be recalled imperfectly, or may never have been completely mastered during initial acquisition. Students who can accurately distinguish between what they know and do not know should be at an advantage while working in such domains, since they are more likely to review and try to relearn imperfectly mastered materials needed for particular tasks, compared with those who are less accurate in estimating their own knowledge. In view of the fact that the knowledge monitoring scores reported in the study, i.e., math and word knowledge, shared about 25% of the variance between them, it would be useful to develop a KMA procedure in the science domain to determine its relationship to achievement in science and engineering.

Several factors are likely to have reduced the magnitude of the effects and the generalizability of the results to other groups of college students. First, many of the students in this sample took less than a full-time schedule of courses. That fact is likely to have decreased the reliability of the gpa, because it was based on fewer courses and credits



than is usually the case after a year of college. Second, it is well known that college grades are often unreliable (Werts, Linn & Jöreskög, 1978; Willingham, Lewis, Morgan, & Ramist, 1990), reducing the magnitude of any correlations with them. The reliability of the grades may have been reduced further by three factors: a) students took dissimilar courses; b) when similar courses were taken they were often taught by different instructors; and c) the differences in students' major fields of study. Criteria other than summary gpa's need to be considered in future research studies.

Further research is also needed to determine the relationships between the KMA procedure and self-report measures of metacognition, study skills, and self-regulated learning strategies measured by the LASSI. Since these constructs are theoretically related to measures obtained with the KMA, a positive empirical relationship between the KMA scores and the LASSI subscales was expected. Surprisingly, none were found in this sample of college students. It is not clear whether these findings would be supported if the LASSI were administered earlier in the students' academic careers, before they experienced a full measure or academic success or failure. We suspect there may be issues related to the method variance, i.e., self-report versus objective scoring, in these two assessment techniques. Finally, the relationship between knowledge monitoring ability, learning and study strategies, and measures of intelligence should be investigated. Sternberg (1991) has suggested that metacognition should be a component of intelligence tests; presumably those who consider metacognition an executive process (Borkowski, Chan, and Muthukrishna, in press) would also agree with that recommendation. Research findings (Schraw, in press) indicate that academically able students have higher knowledge monitoring ability than those less able. Therefore, positive relationships between the KMA procedure and measures of general intellectual ability may be expected. It remains for further research to explore that possibility.



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### APPENDIX A

MULTIPLE REGRESSION

Listwise Deletion of Missing Data

Equation Number 1 Dependent Variable.. KMA Verbal

Block Number 1. Method: Forward Criterion PIN .0500

VCONFIDENCE SFT

Variable(s) Entered on Step Number

VCONFIDENCE voc test conf mean

Multiple R .42060 R Square. .17690 .16942 Adjusted R Square Standard Error .28098

Analysis of Variance

DF Sum of Squares Mean Square Regression 1 1.86648 1.86648 .07895 Residual 110 8.68433

23.64170 Signif F = .0000

----- Variables in the Equation ------

SE B T Sig T Variable Beta .033116 .420599 VCONFIDENCE .161018 4.862 .0000 (Constant) -.263157 .124218 -2.119 .0364

----- Variables not in the Equation -----

Beta In Partial Min Toler T Sig T -.191012 -.210520 .999803 -2.248 .0266

Equation Number 1 Dependent Variable.. KMA Verbal

Variable(s) Entered on Step Number

2.. SFT

Multiple R .46193 R Square .21338 .19895 Adjusted R Square Standard Error .27594

Analysis of Variance

DF . Sum of Squares Mean Square Regression 2 2.25135 1.12568 Residual 109 8.29946 .07614

F =14.78395 Signif F = .0000



------ Variables in the Equation ------Beta T Sig T B SE B Variable .162045 .032525 .423283 4.982 .0000 VCONFIDENCE -2.248 .0266 -.010484 .004663 -.191012 .173697 .014838 .085 .9321 (Constant)

End Block Number 1 All requested variables entered.

Equation Number 1 Dependent Variable.. KMA Math

Block Number 1. Method: Forward Criterion PIN .0500 MCONFIDENCE MOT TMT SFT

Variable(s) Entered on Step Number
1.. MCONFIDENCE math test conf mean

Multiple R .53180 R Square .28281 Adjusted R Square .27652 Standard Error .34784

Analysis of Variance

 DF
 Sum of Squares
 Mean Square

 Regression
 1
 5.43915
 5.43915

 Residual
 114
 13.79341
 .12099

F = 44.95356 Signif F = .0000

------ Variables in the Equation ------

Variable B SE B Beta T Sig T

MCONFIDENCE .294806 .043970 .531798 6.705 .0000 (Constant) -1.028385 .174755 -5.885 .0000

----- Variables not in the Equation ------

 Variable
 Beta In Partial Min Toler
 T Sig T

 MOT
 -.219367 -.258839
 .998507
 -2.849
 .0052

 TMT
 -.203835 -.240406
 .997626
 -2.633
 .0097

 SFT
 -.180162 -.212309
 .995966
 -2.310
 .0227



Equation Number 1 Dependent Variable.. KMA Math

Variable(s) Entered on Step Number

2.. MOT

Multiple R .57520 R Square .33086 Adjusted R Square .31902 Standard Error .33747

Analysis of Variance

 DF
 Sum of Squares
 Mean Square

 Regression
 2
 6.36327
 3.18164

 Residual
 113
 12.86929
 .11389

F = 27.93667 Signif F = .0000

------ Variables in the Equation ------

Beta T Sig T SE B Variable .042691 .290108 .523322 6.796 .0000 MCONFIDENCE .005730 -.219367 -2.849 .0052 -.016322 -1.957 .0529 -.494490 .252732 (Constant)

------ Variables not in the Equation -------

Variable Beta In Partial Min Toler T Sig T

TMT -.115809 -.115549 .666135 -1.231 .2209

SFT -.090131 -.093825 .725118 -.997 .3207

End Block Number 1 PIN = .050 Limits reached.

Listwise Deletion of Missing Data

Equation Number 1 Dependent Variable.. TOTGPA total gpa

Block Number 1. Method: Forward Criterion PIN .0500 KMAV KMAM TST ATT MOT SMI

Variable(s) Entered on Step Number

1.. TST

Multiple R .39284 R Square .15432 Adjusted R Square .14627 Standard Error .70451

Analysis of Variance

 DF
 Sum of Squares
 Mean Square

 Regression
 1
 9.50993
 9.50993

 Residual
 105
 52.11437
 .49633



F =	19.16059	Signi	f F =	.0000			
	v	ariables	in the	Equation	on		
Variable		В	SE B	E	seta	т	Sig T
TST (Constan	.049 t) 1.365		011387 341756	.392	837	4.377 3.996	.0000
	Variab	les not i	n the E	Equation			
Variable	Beta In	Partial	Min 7	Coler	т	Sig T	
KMAV KMAM ATT MOT SMI	.192832 .114426 .069448	.307587 .209263 .091456 .059754 073817	.99 .54 .62	74846 95939 10237 16066 14469	2.182 .937	.0013 .0313 .3511 .5429	
Equation	Number 1	Dependen	t Varia	ble	TOTGPA	tota	l gpa
	(s) Entered KMAV	on Step N	umber				
R Square Adjusted		.48408 .23433 .21961 .67357					
Analysis	of Variance						
Regressi Residual	on	2	14	quares .44046 .18384		n Squar 7.2202 .4536	.3
F =	15.91442	Signi	f F =	.0000			
	v	ariables :	in the	Equation	n		
Variable		В	SE B	В	eta	T	Sig T
KMAV TST (Constant	.697 .044 t) 1.297	078 .0	211663 011026 327402	.286 .347	401	3.297 3.998 3.963	.0013 .0001 .0001
	Variab	les not in	n the E	quation			
Variable	Beta In	Partial	Min T	oler	Т	Sig T	
KMAM		.067734		2381	.689	.4924	
ATT MOT	.133450	.111964		8161 700 <i>6</i>	1.143		
SMI		070718	_	7886 7584	1.333 720	.1854 .4735	



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TOTGPA - Total grade point average

														SONF		SNF FNS	
	ANX	ATT	8	<u>2</u>	MOT	SFT	SM	STA	TST	TMT	KWAM	KMAV	CONF MTH EST	MTH TEST	CONF VOC VOC EST TEST	VOC TEST	Tot A
ANX	1.0000																
ATT .	.5498**	1.0000											•				
8 8	.5904**	5904** 6017** 1.0000	1.0000														
2	0.0941	.2786**	.2786** .2571** 1.0000	1.0000													
MOT	.3165**	.6627**	.6627** .6428**	.4105**	1.0000												
絽	0.0514	.2988.	.2988** .4330** .6576**	.6576**	.5215**	1.0000											
SM	.4458**	.5614**	.4458** .5614** .6528** .4015** .5949**	.4015**	.5949**	.4292**	1.0000										
STA	-0.053	0.136	.3116** .4612** .2891**	.4612**	.2891**	.6273**	.2025	1.0000									
TST	.5788**	.6724**	5788** .6724** .7045** .3149**	.3149**	.6011**	.3427**	.7416**	.1837	1.0000								
TMT	.2747**	.4481**	.2747** ,4481** ,6083** ,2852** ,5737**	.2652**	.5737.	.5443**	.4542**	.4074.	.3993** 1.0000	1.0000							
KMAM	0.1636	-0.107	0.1636 -0.107 -0.159 -0.053		2396**	213	2* -0.097	-0.091	0.025	2293* 1.0000	1.0000	•					
KMAV	0.1189	0.0573	0.0573 -0.002 -0.059		-0.0561	.1902 0.0975		-0.075	0.1445	-0.109	.4908** 1.0000	1.0000					
CONF MTH EST	.2047*	0.1147	.2047* 0.1147 0.0967 0.1615 0.1691	0.1615		0.0709	0.134	.2549**	.2549** 0.0715 0.1012		.3199** .2428** 1.0000	.2428**	1.0000				
CONF MTH TEST	.2308	-0.019	0.0187	-0.021	-0.019 0.0187 -0.021 -0.0417 -0.068		0.0714	0.0714 0.0715 0.0673	0.0673	-0.05	.5309** 0.1673		.5422**	1.0000			-
CONF VOC EST	0.1403	0.1403 0.1242 -0.026	-0.026	0.1306 0.0533		-0.069	0.0286	0.0368	0.0286 0.0368 0.0307 -0.029		0.1639	.3042**	.2653**	0.073	1.0000		
CONF VOC	.3111**	.2776	.3111**   .2776**   .2561**   -0.017	-0.017	0.1032	0.0145	.2047* 0.1048		.3026** 0.1213 0.1704	0.1213	0.1704	.4204** 0.1768	0.1768	0.164	.4656**	1.0000	
TOTGPA	.2294*	.2901**	.2294*  .2901** .2278*  0.1081	0.1081	.2608**	0.1408	.2495**	.2149*	.3754** 1876*		.2382*	.3467** 0.1135	0.1135	0.123	0.0905	0.144 1.0000	1.0000

1. Zero-Order Correlations of LASSI Subscales, Knowledge Monitoring Measures, Confidence Estimates, and GPA

ANX - Anxiety and worry about school performance Legend:

ATT - Attitude

CON - Concentration and attention

INP - Information processing, acquiring knowledge and reasoning

MOT - Motivation

SFT - Self testing, reviewing

SMI - Selecting main ideas and recognizing important information

STA - Use of support techniques and materials

TST - Test strategies and preparing for tests

KMAM - Knowledge monitoring accuracy for mathematics CONF VOC TEST - Confidence in vocabulary test answer KMAV - Knowledge monitoring accuracy for vocabulary CONF VOC EST - Confidence in vocabulary estimates CONF MTH TEST - Confidence in math test answer CONF MTH EST - Confidence in math estimates TMT - Use of time management principles

"." is printed if a coefficient cannot be computed

(2-tailed)

\*\* - Signif. < .01

• - Signif. < .05



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